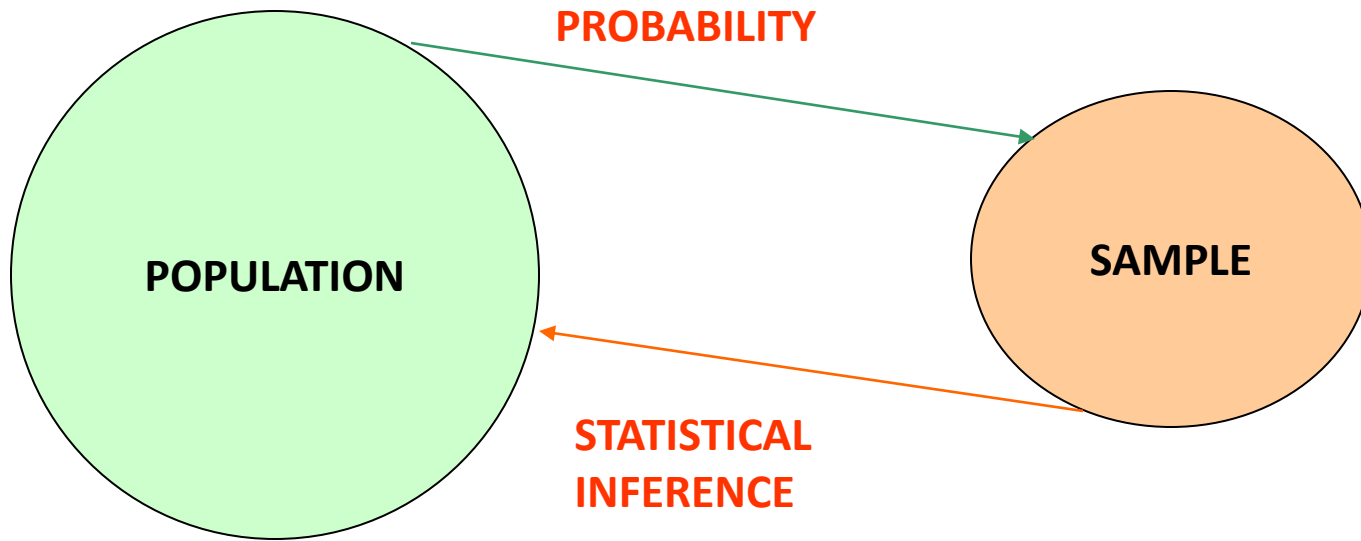


IAM 530
ELEMENTS OF PROBABILITY AND
STATISTICS

LECTURE 2-INTRODUCTION TO
PROBABILITY



WHY PROBABILITY IS NEEDED?

- Nothing in life is certain. We can quantify the uncertainty using **PROBABILITY**
- A probability provides a quantitative description of the chances or likelihoods associated with various outcomes.
- It provides a bridge between descriptive and inferential statistics.

WHAT IS PROBABILITY?

- **CLASSICAL INTERPRETATION**(Frequency of Occurrence)

If a random experiment is repeated, the relative frequency for any given outcome is the probability of this outcome.

Probability of an event: Relative frequency of the occurrence of the event in the long run.

- Example: Probability of observing a head in a fair coin toss is 0.5 (if coin is tossed long enough).

- **SUBJECTIVE INTERPRETATION**(Indication of Uncertainty)

The assignment of probabilities to event of interest is subjective.

- Example: I am guessing there is 50% chance of raining today.

BASIC CONCEPTS

- **Experiment:** is the process by which an observation (or measurement) is obtained.
- **Random experiment:** involves obtaining observations of some kind
- **Population:** Set of all possible observations.
- **Elementary event (simple event):** one possible outcome of an experiment
- **Event (Compound event):** one or more possible outcomes of a random experiment
- **Sample space:** the set of all outcomes for an experiment.

EXAMPLES OF A RANDOM EXPERIMENT

Experiment	Outcomes
Flip a coin	Heads and Tails
Record a statistics test marks	Numbers between 0 and 100
Measure the time to assemble a computer	Numbers from zero and above

SAMPLE SPACE

Countable



Finite number
of elements

Uncountable
(Continuous)



Infinite number of
elements

EXAMPLES

- Countable sample space examples:
 - Tossing a coin experiment
 $S : \{\text{Head, Tail}\}$
 - Rolling a dice experiment
 $S : \{1, 2, 3, 4, 5, 6\}$
 - Determination of the sex of a newborn child
 $S : \{\text{girl, boy}\}$
- Uncountable sample space examples:
 - Life time of a light bulb
 $S : [0, \infty)$
 - Closing daily prices of a stock
 $S : [0, \infty)$

ASSIGNING PROBABILITIES

– Given a sample space $S = \{E_1, E_2, \dots, E_k\}$, the following characteristics for the probability $P(E_i)$ of the simple event E_i *must hold*:

1. $0 \leq P(E_i) \leq 1$ *for each i*

2. $\sum_{i=1}^k P(E_i) = 1$

- The probability of an event A is equal to the sum of the probabilities of the simple events contained in A
- If the simple events in an experiment are **equally likely**, you can calculate

$$P(A) = \frac{\text{total outcomes in } A}{\text{total outcomes in } S}$$

EXAMPLE:

An urn contains 11 red balls and 3 white balls. Two balls are drawn. How likely is that both balls are red?

Suppose we model each ball by positive integer 1-11 for red balls, 12-14 for white balls.

S (the sample space): $\{(1,2),(1,3),\dots,(13,14)\}$

- The set S models all possible draws.
- $14 \times 14 - 14 = 14 \times 13 = 182$ possibilities

E : event of two red balls.

$$E = \{(1,2), (1,3), \dots, (10,11)\}$$

- Number of elements of E = $11 * 11 - 11 = 110$

P: the probability of a finite set is the sum of all its elements

$$P(E) = \sum_{s \in E} P(s) = \sum_{s \in E} \frac{1}{182} = \frac{110}{182}$$

COUNTING RULES

- While forming the sample space at some point, we have to stop listing and to use some counting rules.
- Methods to determine how many subsets can be obtained from a set of objects.

THE $m \cdot n$ RULE

- If an experiment is performed in two stages, with m ways to accomplish the first stage and n ways to accomplish the second stage, then there are mn ways to accomplish the experiment.
- This rule is easily extended to k stages, with the number of ways equal to

$$n_1 n_2 n_3 \dots n_k$$

- **Example:** Toss two coins. The total number of simple events is: $2*2=4$
- **Example:** Toss three dice. The total number of simple events is: $6*6*6=216$
- **Example:** Two balls are drawn from a dish containing two red and two blue balls. The total number of simple events is: $4*3=12$

THE FACTORIAL

- number of ways in which objects can be permuted.

$$n! = n(n-1)(n-2)\dots 2.1$$

$$0! = 1, 1! = 1$$

Example: Possible permutations of $\{1,2,3\}$ are $\{1,2,3\}$, $\{1,3,2\}$, $\{3,1,2\}$, $\{2,1,3\}$, $\{2,3,1\}$, $\{3,2,1\}$. So, there are $3!=6$ different permutations.

PARTITION RULE

- There exists a single set of N distinctly different elements which is partitioned into k sets; the first set containing n_1 elements, ..., the k -th set containing n_k elements. The number of different partitions is

$$\frac{N!}{n_1!n_2!\cdots n_k!} \text{ where } N = n_1 + n_2 + \cdots + n_k.$$

- **Example:** Let's partition $\{1,2,3\}$ into two sets; first with 1 element, second with 2 elements.

Partition 1: $\{1\} \{2,3\}$

Partition 2: $\{2\} \{1,3\}$

Partition 3: $\{3\} \{1,2\}$

$3!/(1! 2!)=3$ different partitions

Example: How many different arrangements can be made of the letters “statistics”?

- $N=10$, $n_1=3$ s, $n_2=3$ t, $n_3=1$ a, $n_4=2$ i, $n_5=1$ c

$$\frac{10!}{3!3!1!2!1!} = 50400$$

PERMUTATIONS

- Any ordered sequence of r objects taken from a set of n distinct objects is called a **permutation** of size r of the objects.

$$P_r^n = \frac{n!}{(n-r)!}$$

where $n! = n(n-1)(n-2)\dots(2)(1)$ and $0! \equiv 1$.

- **Example:** How many 3-digit lock combinations can we make from the numbers 1, 2, 3, and 4?

$$P_3^4 = \frac{4!}{1!} = 4(3)(2) = 24$$

COMBINATION

- Given a set of n distinct objects, any unordered subset of size r of the objects is called a **combination**.

$$C_r^n = \frac{n!}{r!(n-r)!}$$

- **Example:** Three members of a 5-person committee must be chosen to form a subcommittee. How many different subcommittees could be formed?

$$C_3^5 = \frac{5!}{3!(5-3)!} = \frac{5(4)(3)(2)1}{3(2)(1)(2)1} = \frac{5(4)}{(2)1} = 10$$

COUNTING

	Number of possible arrangements of size r from n objects	
	Without Replacement	With Replacement
Ordered	$\frac{n!}{n-r!}$	n^r
Unordered	$\binom{n}{r}$	$\binom{n+r-1}{r}$

ELEMENTARY SET OPERATIONS

INTERSECTION

- The intersection of event A and B is the event that occurs when both A and B occur. It is denoted by $A \cap B$.
- The joint probability of A and B is the probability of the intersection of A and B, which is denoted by $P(A \cap B)$.

UNION

- The union event of A and B is the event that occurs when either A or B or both occur. It is denoted by $A \cup B$.

$$P(A \cup B) = P(A) + P(B) - P(A \cap B)$$

COMPLEMENT

- The complement of event A (denoted by A^c) is the event that occurs when event A does not occur.

$$P(A^c) = 1 - P(A)$$

PROBABILITY FUNCTION AXIOMS

Let S denote a non-empty or finite or countably infinite set and let 2^S denote the set of all subsets of S . A real-valued function P defined on 2^S is a *probability function* if

- For any event E , $0 \leq P(E) \leq 1$.
- $P(S) = 1$.
- If E_1, E_2, \dots are pairwise disjoint events, then

$$P\left(\bigcup_{i=1}^{\infty} E_i\right) = \sum_{i=1}^{\infty} P(E_i)$$

THE CALCULUS OF PROBABILITIES

- If P is a probability function and A and B any sets, then
 - a. $P(B \cap A^c) = P(B) - P(A \cap B)$
 - b. If $A \subset B$, then $P(A) \leq P(B)$
 - c. $P(A \cap B) \geq P(A) + P(B) - 1$ (Bonferroni Inequality)

d.
$$P\left(\bigcup_{i=1}^{\infty} A_i\right) \leq \sum_{i=1}^{\infty} P(A_i) \quad \text{for any sets } A_1, A_2, \dots$$

(Boole's Inequality)

MUTUALLY EXCLUSIVE EVENTS

- When two events A and B are mutually exclusive or disjoint, if A and B have no common outcomes.

$$A \cap B = \emptyset \text{ and } P(A \cap B) = 0$$

- The events A_1, A_2, \dots are pairwise mutually exclusive (disjoint), if

$$A_i \cap A_j = \emptyset \text{ for all } i \neq j.$$

EQUALLY LIKELY OUTCOMES

- The same probability is assigned to each simple event in the sample space, S .
- Suppose that $S = \{s_1, \dots, s_N\}$ is a finite sample space. If all the outcomes are equally likely, then $P(\{s_i\}) = 1/N$ for every outcome s_i .

ADDITION RULE

For any two events A and B

$$P(A \cup B) = P(A) + P(B) - P(A \cap B)$$

COMPLEMENT

- We know that for any event **A**:
 - $P(A \cap A^c) = 0$
- Since either **A** or **A^c** must occur,
 $P(A \cup A^c) = 1$
- so that $P(A \cup A^c) = P(A) + P(A^c) = 1$

Then

$$P(A) = 1 - P(A^c)$$

CONDITIONAL PROBABILITY

The probability that A occurs, given that event B has occurred is called the **conditional probability** of A given B and is defined as

$$P(A|B) = \frac{P(A \cap B)}{P(B)} \quad \text{if } P(B) \neq 0$$

$$0 \leq P(A | B) \leq 1$$

$$P(A | B) = 1 - P(A^c | B), \quad P(A | A) = 1$$

$$P(A_1 \cup A_2 | B) = P(A_1 | B) + P(A_2 | B) - P(A_1 \cap A_2 | B)$$

Example: A red die and a blue die are thrown.

A = { the red die scores a 6 }

B = { at least one 6 is obtained on the two dice }

$$P(A) = \frac{6}{36} = \frac{1}{6} \text{ and } P(B) = \frac{11}{36}$$

$$\begin{aligned} P(A|B) &= \frac{P(A \cap B)}{P(B)} \\ &= \frac{P(A)}{P(B)} \\ &= \frac{1/6}{11/36} = \frac{6}{11} \end{aligned}$$

Example: A bowl contains five candies two red and three blue. Randomly select two candies, and define

- A: second candy is red.
- B: first candy is blue.

$$P(A|B) = P(2^{\text{nd}} \text{ red} | 1^{\text{st}} \text{ blue}) = 2/4 = 1/2$$

$$P(A|\text{not } B) = P(2^{\text{nd}} \text{ red} | 1^{\text{st}} \text{ red}) = 1/4$$

INDEPENDENT VS. NON-INDEPENDENT EVENTS

- If A and B are **independent**, then

$$P(A \text{ and } B) = P(A) \times P(B)$$

which means that conditional probability is:

$$P(B | A) = P(A \text{ and } B) / P(A) = P(A)P(B)/P(A) = P(B)$$

- We have a more general multiplication rule for events that are **not independent**:

$$P(A \text{ and } B) = P(B | A) \times P(A)$$

- In particular, we would like to know whether they are **independent**, that is, if the probability of one event is **not affected** by the occurrence of the other event.

Two events A and B are said to be **independent** if

- $P(A | B) = P(A)$
and
- $P(B | A) = P(B)$

Example

Roll two dice

S =all possible pairs $=\{(1,1),(1,2),\dots,(6,6)\}$

- Let A =first roll is 1; B =sum is 7; C =sum is 8 $P(A|B)=?$;
 $P(A|C)=?$

Solution:

- $P(A|B)=P(A \text{ and } B)/P(B)$

$$P(B)=P(\{1,6\} \text{ or } \{2,5\} \text{ or } \{3,4\} \text{ or } \{4,3\} \text{ or } \{5,2\} \text{ or } \{6,1\})$$
$$= 6/36=1/6$$

$$P(A|B)= P(\{1,6\})/(1/6)=1/6 =P(A)$$

A and B are
independent

- $P(A|C) = P(A \text{ and } C) / P(C) = P(\emptyset) / P(C) = 0$

A and C are disjoint

$$P(C) = P(\{2,6\} \text{ or } \{3,5\} \text{ or } \{4,4\} \text{ or } \{5,3\} \text{ or } \{6,2\})$$
$$= 5/36$$

THE MULTIPLICATIVE RULE FOR INTERSECTIONS

- For any two events, A and B, the probability that both A and B occur is

$$\begin{aligned} P(A \cap B) &= P(A) P(B \text{ given that } A \text{ occurred}) \\ &= P(A)P(B|A) \end{aligned}$$

- If the events A and B are independent, then the probability that both A and B occur is

$$P(A \cap B) = P(A) P(B)$$

BAYES' THEOREM

- Suppose you have $P(B | A)$, but need $P(A | B)$.

$$P(A | B) = \frac{P(A \cap B)}{P(B)} = \frac{P(B | A)P(A)}{P(B)} \text{ for } P(B) \neq 0$$

- Can be generalized to more than two events.

THE LAW OF TOTAL PROBABILITY

Let $B_1, B_2, B_3, \dots, B_k$ be **mutually exclusive** and exhaustive events (that is, one and only one must happen). Then the probability of any event A can be written as

$$\begin{aligned} P(A) &= P(A \cap B_1) + P(A \cap B_2) + \dots + P(A \cap B_k) \\ &= P(B_1)P(A|B_1) + P(B_2)P(A|B_2) + \dots + P(B_k)P(A|B_k) \end{aligned}$$

- Suppose that the events $B_1, B_2, B_3, \dots, B_n$ partition the sample space for some experiment and that A is an event defined on S . For any integer, k , such that $1 \leq k \leq n$ we have

$$P(B_k | A) = \frac{P(A | B_k) P(B_k)}{\sum_{j=1}^n P(A | B_j) P(B_j)}$$

Example: You are living in a dorm. One night the fire alarm goes off. How likely is it that there is a fire?

Here H is the event “there is a fire” and E is the event “the fire alarm goes off.”

You want to know $P(H|E)$.

You estimate that all things being equal a fire is unlikely on a given night, setting $P(H)=0.001$ (roughly one fire in three years).

You know that in a typical semester of about 100 days there are about 3 fire alarms (typically false alarms), so you estimate $P(E)=0.03$.

- Finally you guess that it is nearly certain someone would set off the alarm if there really were a fire, so you estimate $P(E|H)=0.98$.

- By Bayes' Rule,

$$P(H|E)=P(H)P(E|H)/P(E)=(0.001)(0.98)/(0.03)=0.033.$$

- **Example:** A drug company has designed a test for a disease. Through extensive testing, the company reports that the test produces only 1% false positive results (i.e., a healthy person tests positive) and only 2% false negative results (i.e., a person with the disease tests negative).
- Let P be the event “someone tests positive,”
- N be the event “someone tests negative,”
- H be the event “someone is healthy,” and
- D be the event “someone has the disease.”
- Then the company is reporting $P(P|H)=0.01$ (or equivalently $P(N|H)=0.99$) and $P(N|D)=0.02$ (or equivalently $P(P|D)=0.98$).

Suppose you test positive for the disease. How likely is it that you in fact have the disease?

It is tempting but incorrect to say 98% since $P(P|D)=0.98$.

But you want to know $P(D|P)$, which may be quite different. It turns out you do not have enough information yet. Oddly enough you must also know $P(D)$, the prevalence of the disease in your population.

Suppose the disease is rare, occurring in only 0.05% of the population. Then applying the second form of Bayes' we get

$$\begin{aligned} P(D | P) &= \frac{P(D)P(P | D)}{P(D)P(P | D) + P(H)P(P | H)} \\ &= \frac{0.0005 * 0.98}{0.0005 * 0.98 + 0.9995 * 0.01} \approx 0.047 = 4.7\% \end{aligned}$$